

Levelling the Debt-Equity Playing Field: Evidence from Belgium

Muhammad Meki*

September 2022

Abstract

I explore whether a Belgian policy that reduced the corporate tax bias towards debt finance was successful in lowering firm leverage. Using putatively exogenous time series variation in the cost of equity financing and a difference-in-differences strategy that includes similar firms from surrounding countries, I find that the policy did lead to increased equity ratios, with some evidence of a shift from short-term debt to long-term debt.

Keywords: Capital structure; corporate taxation; debt bias; equity financing; small and medium-sized enterprises.

JEL classification: G32, H25, H63.

*Department of International Development, University of Oxford, and Oxford Centre for Islamic Studies. muhammad.meki@qeh.ox.ac.uk. I am grateful to Florin Bilbiie (the editor), and to an anonymous associate editor and referee for very helpful comments. I am also grateful to Johannes Abeler, Thorsten Beck, Stefan Dercon, Michael Devereux, Irem Guceri, Tawfiq Hamid, Michael Koelle, Colin Mayer, Adeel Malik, Simon Quinn, Mohammad Al-Ramadani, Chris Roth, Daniela Scur, Deborah Schanz, Martin Simmler, Daan Struyven, Mesut Tastan, Gabriel Ulyssea, Chris Woodruff, and members of the Centre for Business Taxation at the University of Oxford for helpful discussions and feedback. This work is based on one of the chapters from my PhD, for which I received funding from the Economic and Social Research Council (ESRC) and St John's College, University of Oxford. I also acknowledge support from Pembroke College, University of Oxford, where I undertook a post-doctoral research fellowship generously funded by Chris Rokos; I am also very grateful to Nancy Braithwaite and Pramila Krishnan for their support during the fellowship. All errors are my own.

1 Introduction

In most corporate tax systems, firms can reduce their taxable income by the interest they pay to their debt-holders, while payments to equity-holders receive no such preferential treatment. Several recent financial and economic crises have reignited the academic and public policy debate about this debt bias and the resulting impact of high firm leverage.¹ In this paper, I investigate the effectiveness of one tool for reducing this bias. In 2006, the Belgian government implemented the Notional Interest Deduction (NID), which levelled the playing field between equity and debt by allowing firms to reduce their taxable income based on their amount of equity capital. The main motivation for the policy was a European Commission ruling, not domestic economic conditions. This provides putatively exogenous time series variation in the cost of equity financing, which I exploit as part of a difference-in-differences strategy. I use several firm- and country-level controls, and data from firms in countries surrounding Belgium, which are economically integrated and culturally similar, and did not experience major tax policy changes around the time of the NID's implementation. I find that the NID was successful in increasing equity ratios of Belgian firms by approximately three percentage points (compared to a pre-NID mean equity-to-asset ratio of around 39%). There is also some evidence of a shift from short- to long-term debt. Results are robust to including several control variables, to the use of a matching strategy to increase the comparability of the non-Belgian sample, and to several alternative specifications. Analysis of heterogeneous effects provides some evidence that the greatest increase in equity occurred for firms in sectors with the highest pre-NID revenue volatility, which may particularly benefit from the implicit insurance of equity finance.

In contrast to many studies exploring the effects of taxes on firm financing, the NID generated large putatively exogenous time series variation in the cost of equity. The existing literature on the NID has shown mixed results, likely due to a lack of statistical power in studies that investigated the initial impacts, especially those that focused on SMEs. For example, [Van Campenhout and Van Caneghem \(2013\)](#) and [Kestens, Van Cauwenberge, and Christiaens \(2012\)](#) find conflicting evidence for the impact on SMEs (using sample sizes of 614 and 4,474 respectively). [Princen \(2012\)](#), using a sample of 3,332 Belgian firms one year after the NID, argues that the NID led to lower leverage relative to French firms. [Panier, Pérez-González, and Villanueva \(2013\)](#) use a larger sample and do find a significant short-term effect on equity ratios. Other papers focus on the impact on multinationals, often finding evidence of lower leverage achieved through intra-group tax arbitrage strategies ([Hebous & Ruf, 2017](#); [Konings, Lecocq, & Merlevede, 2018](#)).² I use a large and longer-term sample than much of the existing literature, and I show that the NID was effective in reducing leverage when the sample includes large firms that

¹ The Mirrlees Review found that the bias towards debt discourages business investment and exposes firms to greater risk of bankruptcy, the LSE Growth Commission argues that it distorts long-term investment incentives, and The Economist magazine describes it as a “senseless subsidy that wins the title of the world’s worst economic distortion” ([Aghion et al., 2013](#); [Economist, 2015](#); [Mirrlees et al., 2010](#)).

² Other studies have focused on the effect of the NID for financial institutions ([Célérier, Kick, & Ongena, 2018](#); [Schepens, 2016](#))

are part of a broader group (thus providing a replication of previous results) as well as using a sample of only medium-sized, stand-alone firms (for which the evidence to date is scant).

2 Identifying the impact of the NID

‘Allowance for Corporate Equity’ (ACE) systems, of which the Belgian Notional Interest Deduction (NID) is one type, mitigate the classic debt tax bias by introducing a similar deduction for equity.³ Several countries have introduced ACE-like reforms, but many were implemented in a restricted form and for a short period, and often occurred at the same time as other tax policy changes (Klemm, 2007; Zangari, 2014). The main motivation for the implementation of the Belgian NID in 2006 was not domestic economic conditions; the policy was a response to a European Commission ruling on ‘coordination centres’, which were companies that provided financial, accounting and administrative services on behalf of their multinational group, and had benefitted from advantageous tax legislation from 1982 to 2003 (C  lerier et al., 2018; Panier et al., 2013; Valenduc, 2009). In 2003, the European Commission ruled that coordination centres contravened EU rules on state aid; since many such centres were equity financed, the NID offered a way for the tax authorities to mitigate the risk of multinationals leaving Belgium.

The NID allowed firms to reduce their taxable profit using a simple formula: the product of the firm’s adjusted book value of equity and the average rate on 10-year Belgian Government Bonds. If the firm’s rate of return on equity was less than or equal to 3%, it would face a zero corporate tax rate. A higher return on equity would only be subject to corporate tax at a rate equal to the excess of its return over and above the NID rate of 3%, thereby reducing its effective corporate tax rate. All Belgian firms were eligible for the NID, but the policy was implemented in a way that complicated the incentives for small firms, defined as having fewer than 100 employees and not exceeding more than one of the following three thresholds: (i) total assets of €3.65 million; (ii) revenue of €7.3 million; or (iii) 50 employees. Small firms were actually provided with an additional incentive to implement the NID (an extra 0.5% deduction), however, two concurrent policy provisions made it less attractive for smaller firms to increase their equity (Van Campenhout & Van Caneghem, 2013; Zangari, 2014). The first was the elimination of a tax credit for firms that increase their equity capital above the maximum that it had reached in the previous three years. The second was the discontinuation of a programme that allowed firms to reduce their taxable income based on the value of their investments that were funded with retained earnings. Given the combined allowance amount of €38,600, it is possible to calculate a cutoff above which firms would have benefited from the NID. Specifically, assuming a 2.7% NID rate (the average rate since inception, as of 2013), firms with total equity greater than €1.43 million would have benefited from increasing their equity capital through the NID. This is equivalent to total assets of €3.7 million, using

³ ACE systems are based on the concept of a ‘pure profits’ tax that obtains revenues in a non-distorting manner due to the neutral taxation of capital income (Boadway & Bruce, 1984; Devereux & Freeman, 1991). ACE systems have been promoted by the Mirrlees Review and by the Institute for Fiscal Studies as a way of achieving tax neutrality between debt and equity financing (Gammie, 1991; Mirrlees et al., 2010).

the sample average pre-NID equity-to-asset ratio of 38.6%. I then use a value of €4 million in pre-NID total assets as a cutoff to distinguish small firms from larger firms.

I use a difference-in-differences (DiD) strategy to exploit time series variation in the cost of equity arising from the NID. The ‘treatment group’ is Belgian firms with total assets greater than €4 million in the two years prior to the NID’s introduction (henceforth referred to as ‘large’), and the control group is similarly sized firms from surrounding countries (France, Germany, Luxembourg and the Netherlands). Firms in these countries are geographically and culturally close, and their relative economic integration means that they are exposed to similar shocks and trends. Furthermore, there were no major tax policy changes in these countries around the introduction of the NID in 2006 (Konings, Lecocq, & Merlevede, 2018; Panier, Pérez-González, & Villanueva, 2013). Appendix Figure A.1 illustrates the strong correlation between GDP growth of Belgium and surrounding countries. In some specifications, I also include small firms from both Belgium and surrounding countries as a within-country control for domestic time trends (i.e. a ‘triple-difference’ specification), which mitigates concerns about bias arising from leverage ratios trends that are unrelated to the NID. I also draw on the extensive theoretical and empirical corporate finance literature on capital structure determinants when including baseline control variables.⁴

Figure 1 provides a preview of results and visual support for the ‘common trends’ identifying assumption of the DiD estimator: small and large firms from surrounding control countries appear to be on a parallel trend throughout the ten-year period; in contrast, while trends appear to be parallel for small and large Belgian firms up until the implementation of the NID in 2006, after 2006 equity ratios jump up significantly for large Belgian firms only. Appendix Section A.1 contains several regressions that formally test for the parallel trends assumption, using the method described by Kahn-Lang and Lang (2020). In each of the 14 tests, I fail to reject the null of equal trends, whether it is testing for equal trends between large Belgian firms and large non-Belgian firms, or testing for equal trends between small and large firms within each country. Notwithstanding this, in the regressions I explicitly allow for a differential trend in the post-NID period, following recommendations in the DiD literature (Bilinski & Hatfield, 2018).

3 Data and baseline summary statistics

The sample consists of ten years of annual data for firms from Belgium, France, Germany, Luxembourg, and the Netherlands. The data was extracted from Bureau van Dijk’s Orbis database, covering 2004 to 2013. The Belgian data are of a high quality, since Bureau van Dijk uses financial information filed at the National Bank of Belgium, and given that all Belgian firms are legally required to report the total

⁴ The most relevant theory is the ‘trade-off theory’ of capital structure, which highlights the importance of taxes and bankruptcy costs for firms’ choice of leverage (Kraus & Litzenberger, 1973; Miller, 1977; Myers, 1984), and predicts that leverage would decrease as a result of the NID. Other influential theories make predictions about the relationship between leverage and firm size, profitability, tangibility and liquidity ratios (Hall, Hutchinson, & Michaelas, 2004; Myers, 1984; Myers & Majluf, 1984; Rajan & Zingales, 1995), and I include their pre-NID levels as controls, as well as their pre-NID growth rates, following Quinn (2014).

value of their assets, equity and non-equity liabilities. The coverage of the French data is also very comprehensive, given that all firms in France must file standardised financial accounts to the authorities. The coverage is less comprehensive for Germany, Luxembourg, and the Netherlands, where there are relatively less standardised and systematic disclosure requirements for all private firms (Panier, Pérez-González, & Villanueva, 2013; Princen, 2012). I use all available data for limited liability corporations with information available for at least one year in the two years prior to the NID. I exclude firms whose assets were below €500,000 in any pre-NID year, since their leverage decisions may reflect household considerations. Following Kalemlı-Özcan, Laeven, and Moreno (2018) and Barbiero, Popov, and Wolski (2020), I drop firm-year observations in which total assets, fixed assets, short-term debt, long-term debt, or total shareholder funds have negative values, and I also drop country-specific sectors, such as agriculture and mining, sectors with high government ownership, and heavily regulated sectors such as finance. To account for outliers, I winsorize all variables at the 2.5% level in each tail of the distribution.

I use a balanced panel containing up to 1,154,000 firm-year observations from 126,172 firms (comprising 35,406 Belgian firms, with 79% of the non-Belgian data coming from France). Appendix Section A.2 presents baseline summary statistics. The average size of Belgian and French firms is similar, with mean log assets of 7.28 and 7.31 respectively (based on total assets in thousands of €), while firms in other countries are larger on average, which is unsurprising given the aforementioned reporting requirements. The average age of Belgian and French firms is comparable (26.4 and 26.7 years respectively). While there are some baseline differences, these do not invalidate the common trends assumption required for identification of the DiD estimator. Nonetheless, to further increase confidence in the results I follow the recommendation of Ryan, Kontopantelis, Linden, and Burgess Jr (2019) and McKenzie (2020) by implementing a matched DiD using a more comparable set of firms from surrounding countries, matched on the pre-NID profitability, tangibility, and equity ratios, as well as business sector.⁵ In the unmatched sample, mean equity ratios, defined as the sum of shareholders' funds and retained earnings as a proportion of its total assets, are slightly lower for Belgian firms (38.5%) compared to French firms (39.2%). Column 5 of Table A.3 demonstrates that the matching strategy is successful in bringing the characteristics of the matched sample closer to the Belgian sample, in terms of profits to total assets (a mean of 6.7% in Belgium), the tangible to fixed assets ratio (a mean of 31% in Belgium), and the sectoral distribution.

4 Results

I exploit time series variation in the cost of equity financing and a difference-in-differences strategy:

$$y_{it} = \beta_0 + \beta_1 Post_t + \beta_2 Treatment_i + \beta_3 Post_t * Treatment_i + \beta_4 \mathbf{X}_{i0} + \varepsilon_{it} \quad (1)$$

⁵ I implement a nearest neighbour propensity score matching procedure (with replacement), enforcing common support and calipers of 0.01, following Ryan et al. (2019). Subsequent weights (representing the number of Belgian observations for which the particular non-Belgian firm is a match) are used as frequency weights in the subsequent empirical estimation.

where y_{it} is the leverage ratio of firm i at time t (using equity ratios or debt ratios), $Post_t$ is an indicator variable for the post-NID period (2006 onwards), $Treatment_i$ is an indicator for Belgian firms with total assets above €4 million in the pre-NID period, and control firms are those in surrounding countries (France, Germany, Luxembourg, and the Netherlands). The coefficient of interest is β_3 , which will be positive if the NID led to an increase in equity ratios for Belgian firms (or negative when the dependent variable is the debt ratio). In some specifications, I include baseline controls for variables that have been shown to be important determinants of firm leverage. In other specifications, I also include small firms from Belgium and surrounding countries as an additional within-country control group.

Results for the effect of the NID on equity ratios are presented in Table 1. Columns 1, 2, 5, and 6 include only stand-alone firms, with the remaining columns including all firms that are part of a larger group / subsidiary of a multinational. Columns 1 and 5 include firm fixed effects, and columns 2, 3, 6, and 7 present OLS regressions with no controls other than a dummy variable for Belgium (to control for permanent differences between Belgium and the other countries), a dummy for the post-NID period (which controls for common trends), a dummy for large firms (only in the triple-difference specifications that include small firms), and the interaction of all these variables. Columns 4 and 8 include sector fixed effects and the aforementioned baseline controls. Beginning with the first four columns (which include small firms, and contain up to 1,154,684 firm-year observations), results indicate a positive impact of the NID on equity ratios, with a coefficient on $BEL * Post * Large$ that ranges from 2.68 percentage points to 3.66 percentage points, compared to the baseline mean equity ratio of around 39%, and with all coefficients significant at the 1% level. In columns 5 to 8, the sample is restricted to firms with total assets above €4 million in the pre-NID period (i.e. dropping all small firms), but results are similar; the coefficient on $BEL * Post$ indicates a positive impact of the NID on equity ratios that ranges from 2.09 to 3.16, with each coefficient again statistically significant at the 1% level.

In Table 2, I repeat the analysis using the debt ratio as the dependent variable. Results provide reassurance about the quality of the balance sheet data; estimated impacts on net debt (the sum of effects on short-term and long-term debt) are of similar magnitude (and opposite sign) to those on equity ratios, confirming that the post-NID period was associated with a lower leverage ratio for Belgian firms. Moreover, there is evidence of a more nuanced effect on debt ratios, when investigating separately the impact on short- and long-term debt. Specifically, there is strong evidence of the NID leading to lower short-term debt ratios, with mixed evidence of a small positive impact on the long-term debt ratio (suggesting some substitution from short-term to long-term debt). In columns 1 to 3, where small firms are included, and the number of controls are gradually increased, the coefficient on $BEL * Post * Large$ ranges from -4.32 to -5.58, with all coefficients significant at the 1% level. Columns 4 to 6 repeat the analysis without the within-country control for small firms, and find a similar impact, with the coefficient on $BEL * Post$ ranging from -3.06 to -3.78, with all coefficients again significant at the 1% level. This is in contrast to the results in columns 7 and 8, which present the same specification (with the full set of controls) but using the

long-term debt ratio as the dependent variable, and alternatively using the sample with and without small firms. In column 7, the coefficient on $BEL * Post * Large$ is actually *positive*, with a coefficient of +3.58 percentage points; however, the coefficient on $BEL * Post$ in column 8 is insignificantly different from zero.

In Table 3, the analysis from Tables 1 and 2 is replicated using a matched set of comparison firms. The findings are consistent with the previous results; there is a relatively large and statistically significant positive impact of the NID on the equity ratios of Belgian firms. Further, there is an even larger (in absolute magnitude) decrease in short-term debt ratios, with some mixed evidence of a partial offset with greater long-term debt.

5 Robustness and further analysis

In the Appendix, I present results using a number of alternative specifications, to provide further robustness checks for the main results. First, in Appendix Section A.3, I demonstrate that results are robust to collapsing the dataset to a two-period panel (pooling the pre- and post-NID periods). This follows the recommendation of [Bertrand, Duflo, and Mullainathan \(2004\)](#) to address concerns about serial correlation in DiD estimations (especially when the dependent variable is positively serially correlated), which may lead to an underestimation of standard errors. Second, I mitigate potential concerns about survivorship bias stemming from the use of a balanced panel, by repeating all the analysis with the full (unbalanced) panel. Results in Section A.4 demonstrate that the results are robust and that the magnitudes of the estimated coefficients increase in many of the specifications. Third, in Section A.5, I demonstrate robustness of the results to systematically adding France, Germany, Luxembourg, and the Netherlands as the control group (compared to Tables 1, 2, and 3, where all countries are included at once).

Finally, in appendix section A.6, I explore whether the NID had heterogeneous impacts on firms in the most risk-exposed sectors, for whom the performance-contingent nature of equity payments may be particularly beneficial. Using several specifications and sample selection criteria, I find that firms in the most risk-exposed sectors (captured by the sector-level coefficient of variation of revenue volatility in the pre-NID period) experienced the greatest increase in equity ratios as a result of the NID. Further, I show that these also happen to be the sectors that created the most jobs across the EU between 2000 and 2014.

6 Conclusion

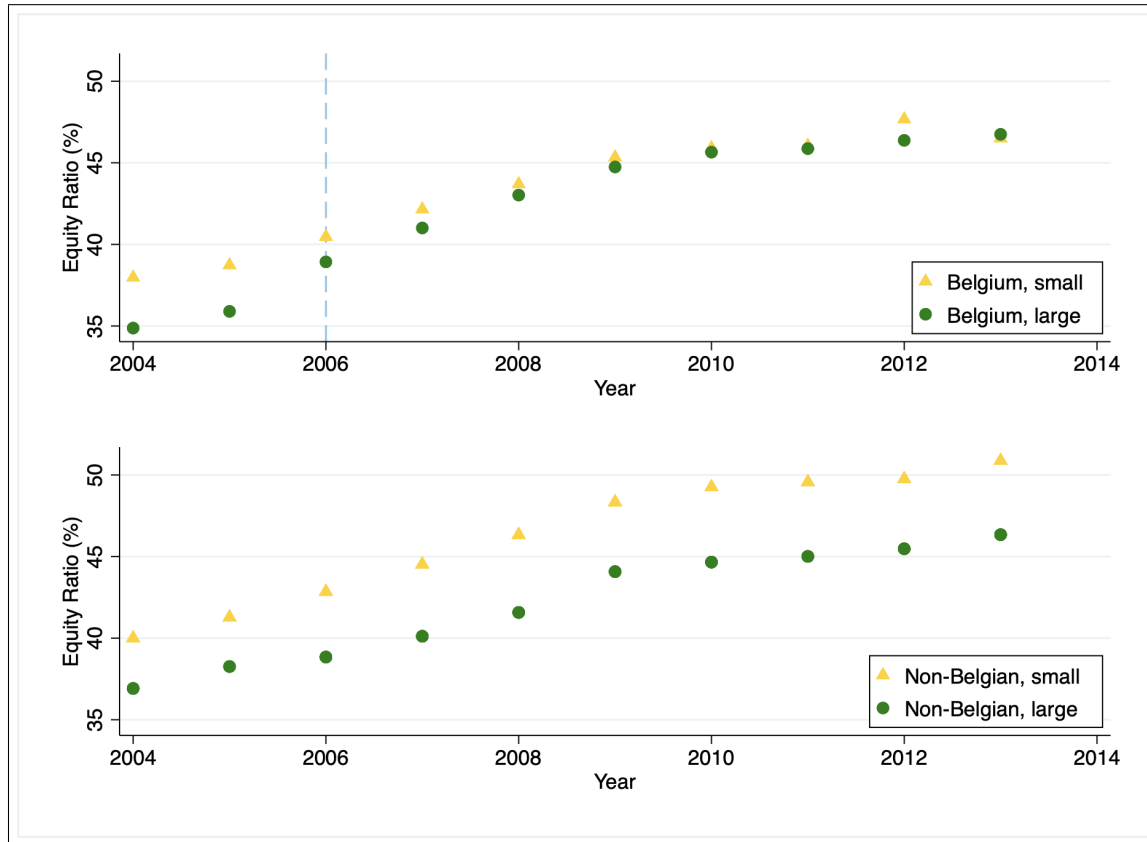
Debt finance serves many critical functions in the economy, allowing cash-constrained firms to borrow against future profits, without diluting ownership and control. However, systematically discriminating against equity can lead to excessive leverage. This can increase financial system fragility, which may be an especially important concern during recoveries from economic and financial crises. For example, in response to COVID-19, governments across the world implemented over 1,600 measures directly aimed

at supporting firms, with a large share of those involving debt finance (Cirera et al., 2021). Equity-like financing can provide firms with a flexible form of capital that spreads risk more effectively; payments to equity-holders can be reduced in times of financial stress, and equity does not need to be re-financed. While there has been much discussion about the need to encourage greater capital buffers for firms, especially following the 2008 global financial crisis and subsequent European debt crisis, there is a tension in the conflicting policy of asking banks to hold more equity through capital requirements while simultaneously encouraging them to hold more debt through the tax system, which significantly complicates and distorts the functioning of the financial system (Mayer, Micossi, Onado, Pagano, & Polo, 2018).

In this paper, I find that a policy to level the playing field between debt and equity was successful in reducing leverage for firms, leading to higher equity ratios and some suggestion of a shift from short-term debt towards longer-term debt. While this highlights the exciting potential of ACE-type models, there is a need to carefully consider the exploitation of policies, for example through intra-group tax arbitrage strategies by multinationals (Hebous & Ruf, 2017). Nonetheless, ACE-like models provide one possible tool to answer the increasing calls for policymakers to take an active role in dealing with high firm leverage, especially for SMEs, who were disproportionately affected by the COVID-19 crisis and may find it harder to issue conventional equity through capital markets (and / or may be more averse to ownership dilution). Other novel proposals include linking loan repayments to business returns, as well as more radical suggestions such as a pan-European Pandemic Equity Fund that allows all European citizens to participate in a risk- and reward-sharing mechanism for firms (Boot et al., 2020; IMF, 2020; OECD, 2020; Revoltella, Maurin, & Pal, 2020; Sandbu, 2020). This is undoubtedly an exciting and important time for further empirical and theoretical work in this area.

Tables and Figures

Figure 1: EVOLUTION OF EQUITY RATIO



Note: This figure displays the average equity-to-asset ratio for firms in every year in the sample. The top panel displays the equity ratio for Belgian firms, categorised by size: (i) large firms (those with total assets greater than €4 million in the two years prior to the implementation of the NID in 2006, indicated by the blue dashed line) who comprise the ‘treatment group’ for whom the NID provided a strong incentive to increase equity ratios; (ii) small firms, those with assets less than €4 million in the two pre-NID years, who (due to the nuances of policy implementation, discussed in Section 2) form a within-country control for the treatment group. The bottom panel displays the evolution of equity ratios for all non-Belgian firms (from France, Germany, Luxembourg, and the Netherlands), with a similar categorisation by size. Format tests of parallel trends, using a number of specifications, are provided in Appendix Section A.1

Table 1: EFFECT OF THE NID ON EQUITY RATIOS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Belgium * Post * Large	3.65*** (0.611)	3.66*** (0.611)	3.00*** (0.259)	2.68*** (0.250)				
Belgium * Post	-0.67*** (0.128)	-0.67*** (0.128)	0.16 (0.099)	0.14 (0.099)	2.98*** (0.597)	2.99*** (0.597)	3.16*** (0.240)	2.09*** (0.225)
Belgium * Large		0.08 (1.114)	0.82** (0.411)	-2.04*** (0.394)				
Post * Large	-1.36*** (0.348)	-1.38*** (0.348)	-2.31*** (0.113)	-1.89*** (0.120)				
Belgium		-2.29*** (0.217)	-0.25 (0.160)	2.05*** (0.162)		-2.21** (1.093)	0.57 (0.379)	0.04 (0.377)
Large		-3.05*** (0.662)	-2.13*** (0.193)	0.45* (0.275)				
Post	7.04*** (0.081)	7.05*** (0.081)	5.93*** (0.051)	-0.48*** (0.065)	5.68*** (0.338)	5.67*** (0.338)	3.62*** (0.100)	-1.48*** (0.130)
Dependent variable mean	39.52	39.52	38.95	38.95	36.73	39.52	38.95	37.38
Small firms	✓	✓	✓	✓				
Group / subsidiary firms			✓	✓			✓	✓
Firm fixed effects	✓				✓			
Sector-year fixed effects				✓				✓
Baseline controls				✓				✓
Observations	479,636	479,636	1,154,684	1,087,382	22,799	22,799	205,966	181,737

Note: This table presents results for the impact of the NID on firm leverage. The dependent variable in all columns is the equity to total assets ratio. Post and Large are indicator variables for the post-NID period and for having greater than €4 million in pre-NID total assets (the cut-off above which Belgian firms would have benefited from the NID). Baseline controls consist of the pre-NID levels and growth rates of the firm's profitability, tangibility and liquidity ratios, the pre-NID growth rate of the equity ratio itself, as well as country-level controls for GDP and statutory corporate tax rates (top federal-level rate). Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

Table 2: EFFECT OF THE NID ON DEBT RATIOS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	S/T	S/T	S/T	S/T	S/T	S/T	L/T	L/T
	debt	debt	debt	debt	debt	debt	debt	debt
Belgium * Post * Large	-5.31*** (0.653)	-4.32*** (0.260)	-5.58*** (0.261)				3.58*** (0.189)	
Belgium * Post	1.53*** (0.125)	1.26*** (0.097)	1.28*** (0.097)	-3.78*** (0.641)	-3.06*** (0.241)	-3.38*** (0.238)	-4.24*** (0.078)	-0.26 (0.169)
Belgium * Large	13.11*** (1.166)	9.17*** (0.426)	6.27*** (0.400)				-3.75*** (0.242)	
Post * Large	1.85*** (0.404)	2.44*** (0.127)	3.45*** (0.142)				-1.76*** (0.106)	
Belgium	-7.40*** (0.208)	-8.80*** (0.155)	-6.87*** (0.160)	5.71*** (1.147)	0.37 (0.396)	-1.18*** (0.387)	7.96*** (0.100)	4.67*** (0.226)
Large	-2.86*** (0.719)	-1.13*** (0.213)	-1.85*** (0.276)				1.51*** (0.148)	
Post	-4.58*** (0.078)	-4.27*** (0.052)	-1.77*** (0.092)	-2.73*** (0.397)	-1.83*** (0.116)	-1.57*** (0.189)	-0.15* (0.078)	0.58*** (0.155)
Dependent variable mean	44.14	46.81	46.81	44.14	46.81	48.26	9.29	7.82
Small firms	✓	✓	✓				✓	
Group / subsidiary firms		✓	✓		✓	✓	✓	✓
Sector-year fixed effects			✓			✓	✓	✓
Baseline controls			✓			✓	✓	✓
Observations	479,454	1,154,355	1,087,177	22,772	205,890	181,699	1,087,086	181,672

Note: This table presents results for the impact of the NID on firm leverage, from the perspective of debt ratios. The dependent variable in columns 1 to 6 is the short-term debt to total assets ratio and in columns 7 and 8 it is the ratio of long-term debt to total assets. Post and Large are indicator variables for the post-NID period and for having greater than €4 million in pre-NID total assets (the cut-off above which Belgian firms would have benefited from the NID). Baseline controls consist of the pre-NID levels and growth rates of the firm's profitability, tangibility and liquidity ratios, the pre-NID growth rate of the equity ratio itself, as well as country-level controls for GDP and statutory corporate tax rates (top federal-level rate). Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

Table 3: ESTIMATION WITH MATCHED SAMPLE

	(1)	(2)	(3)	(4)	(5)	(6)
	Equity	S/T debt	L/T debt	Equity	S/T debt	L/T debt
Belgium * Post * Large	2.10*** (0.410)	-4.54*** (0.463)	2.99*** (0.368)			
Belgium * Post	0.22 (0.168)	-0.22 (0.174)	-3.86*** (0.148)	1.82*** (0.315)	-3.57*** (0.383)	-0.59* (0.307)
Belgium * Large	-3.46*** (0.665)	7.45*** (0.648)	-3.46*** (0.446)			
Post * Large	-1.44*** (0.349)	3.88*** (0.410)	-2.72*** (0.336)			
Belgium	-0.32 (0.277)	-3.41*** (0.254)	8.23*** (0.176)	-3.02*** (0.606)	4.09*** (0.624)	4.34*** (0.381)
Large	1.94*** (0.671)	-4.32*** (0.634)	1.93*** (0.447)			
Post	-0.16 (0.155)	-0.57*** (0.194)	1.63*** (0.168)	-1.08*** (0.289)	0.29 (0.387)	0.88*** (0.306)
Dependent variable mean	39.40	42.09	14.16	39.40	42.09	14.16
Small firms	✓	✓	✓			
Group / subsidiary firms	✓	✓	✓	✓	✓	✓
Sector-year fixed effects	✓	✓	✓	✓	✓	✓
Baseline controls	✓	✓	✓	✓	✓	✓
Observations	481,575	481,525	481,519	78,045	78,033	78,028

Note: This table replicates the main analysis from Tables 1 and 2, using a matched sample of firms from surrounding countries to increase comparability with the Belgian sample. Matching is performed using baseline (pre-NID) profitability ratios, tangibility ratios, equity ratios, and business sector. A nearest neighbour propensity score matching procedure (with replacement) is implemented, enforcing common support and calipers of 0.01. Each specification also contains the same set of baseline control variables as in Tables 1 and 2. Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

Appendix

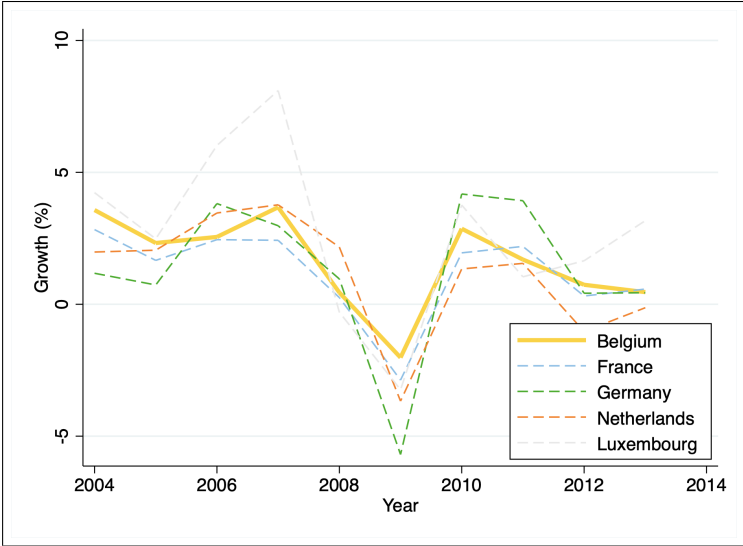


Figure A.1: GDP GROWTH

The above graph illustrates the similarity in the growth rate of GDP for Belgium and surrounding countries between 2004 and 2013. The correlation coefficient between Belgium and France is 0.95, and it is 0.80, 0.85, and 0.87 between Belgium and Germany, the Netherlands, and Luxembourg, respectively. Data source: World Development Indicators, World Bank.

A.1 Testing for parallel trends

Table A.1: PARALLEL TRENDS TEST

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Large * 2005	1.02*** (0.257)	1.02*** (0.257)	1.34*** (0.191)	1.27*** (0.193)	1.34*** (0.191)	1.35*** (0.191)	1.05*** (0.214)	1.04*** (0.215)
Small * 2005	0.76*** (0.075)	0.76*** (0.075)	1.27*** (0.055)	1.27*** (0.056)	1.27*** (0.055)	1.27*** (0.055)	0.87*** (0.067)	0.87*** (0.068)
Large	-3.10*** (0.899)	-2.66*** (0.885)	-3.08*** (0.668)	-1.87** (0.767)	-3.08*** (0.668)	-3.21*** (0.665)	-3.07*** (0.814)	-3.09*** (0.804)
Sample	BEL	BEL	FRA	FRA	ALL	ALL	MATCH	MATCH
Sector fixed effects		✓		✓		✓		✓
Test: equal trend	0.320	0.340	0.725	0.996	0.725	0.710	0.420	0.436
Observations	40,950	40,900	55,006	52,556	55,006	55,000	51,210	51,164

This table presents regressions that formally test the parallel trends assumption, the p -value for which is displayed in the penultimate row of the table. Columns 1 and 2 present regressions using only Belgian firms, in columns 3 and 4 it is French firms, in columns 5 and 6 it is all non-Belgian firms (from France, Germany, Luxembourg and the Netherlands), and in columns 7 and 8 it is the sample of matched firms. Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** $p < 0.001$, ** $p < 0.05$, * $p < 0.10$.

Table A.2: PARALLEL TRENDS, EXCLUDING SMALL FIRMS

	(1)	(2)	(3)	(4)	(5)	(6)
Belgium * 2005	1.02*** (0.257)	1.02*** (0.259)	1.02*** (0.257)	1.02*** (0.259)	1.03*** (0.259)	1.03*** (0.262)
Control * 2005	1.27*** (0.193)	1.27*** (0.194)	1.34*** (0.191)	1.35*** (0.192)	1.09*** (0.381)	1.09*** (0.384)
Belgium	-3.65*** (1.166)	-0.80 (1.163)	-2.05* (1.098)	0.50 (1.085)	-2.25 (1.901)	-1.02 (1.821)
Non-Belgian sample	FRA	FRA	ALL	ALL	MATCH	MATCH
Sector fixed effects		✓		✓		✓
Test: equal trend	0.447	0.436	0.319	0.305	0.909	0.897
Observations	3,854	3,852	4,562	4,556	2,508	2,506

This table presents regressions that formally test the parallel trends assumption, the p -value for which is displayed in the penultimate row of the table, and excluding all small firms (i.e. just including large firms in Belgium and control countries). Columns 1 and 2 present regressions using French firms as control firms, in columns 3 and 4 firms from all countries surrounding Belgium are included (France, Germany, Luxembourg and the Netherlands), and in columns 5 and 6 it is the sample of matched firms. Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** $p < 0.001$, ** $p < 0.05$, * $p < 0.10$.

A.2 Summary statistics

Table A.3: BASELINE SUMMARY STATISTICS

	(1)	(2)	(3)	(4)	(5)	(6)
	Belgium	France	Germany	Netherlands	Luxembourg	Matched
Equity ratio	38.48 (24.66)	39.15 (21.25)	30.01 (21.45)	37.42 (22.82)	36.81 (22.25)	40.87 (21.74)
Short-term debt ratio	42.03 (24.05)	50.99 (21.31)	26.06 (27.54)	50.60 (23.78)	48.42 (24.48)	46.00 (21.16)
Profitability ratio	6.73 (9.88)	9.87 (10.91)	8.51 (11.19)	11.62 (12.72)	8.32 (11.48)	8.58 (10.45)
Log assets	7.28 (1.14)	7.31 (1.17)	8.24 (1.62)	9.07 (1.62)	8.79 (1.87)	7.35 (1.16)
Tangibility ratio	30.93 (25.59)	13.31 (15.54)	22.99 (21.79)	17.79 (20.27)	15.52 (17.34)	25.81 (24.22)
Firm age	26.43 (13.34)	26.71 (12.96)	34.01 (30.05)	39.17 (27.36)	25.36 (18.17)	27.93 (13.47)
Sector: Services	0.22 (0.41)	0.24 (0.43)	0.21 (0.40)	0.18 (0.39)	0.16 (0.37)	0.24 (0.43)
Sector: Wholesale trade	0.29 (0.45)	0.22 (0.42)	0.21 (0.41)	0.36 (0.48)	0.27 (0.45)	0.30 (0.46)
Sector: Retail trade	0.13 (0.34)	0.15 (0.36)	0.08 (0.27)	0.04 (0.19)	0.10 (0.29)	0.12 (0.33)
Sector: Construction	0.17 (0.37)	0.17 (0.37)	0.15 (0.36)	0.14 (0.35)	0.16 (0.37)	0.16 (0.37)
Sector: Manufacturing	0.19 (0.39)	0.21 (0.40)	0.35 (0.48)	0.27 (0.44)	0.20 (0.40)	0.18 (0.38)
Number of firms	35,406	71,803	5,596	2,532	209	10,626

Note: This table presents baseline summary statistics (using the first available year in the dataset, 2004). Log assets are based on total assets in €thousands. All ratios are expressed as a percentage of total assets. Equity is calculated as the sum of shareholder capital and retained earnings, while the short-term debt ratio uses the firm's current liabilities. The profitability ratio uses firm profits before taxes and the tangibility ratio uses net fixed assets. Columns 1 to 5 summarise data from all firms in Belgium, France, Germany, the Netherlands, and Luxembourg, respectively, while column 6 summarises the matched sample of firms obtained using the matching method described in Section 2.

A.3 Estimation with a two-period panel

Table A.4: TWO-PERIOD PANEL

	(1)	(2)	(3)	(4)	(5)	(6)
	Equity	Equity	S/T debt	S/T debt	L/T debt	L/T debt
Belgium * Post * Large	3.23*** (0.615)	3.26*** (0.624)	-5.12*** (0.661)	-4.92*** (0.660)	2.72*** (0.457)	2.57*** (0.457)
Belgium * Post	-0.64*** (0.128)	-0.65*** (0.130)	1.53*** (0.125)	1.56*** (0.125)	-4.37*** (0.102)	-4.39*** (0.103)
Belgium * Large	0.30 (1.144)	-0.51 (1.000)	12.93*** (1.191)	8.23*** (0.992)	-9.94*** (0.691)	-5.52*** (0.606)
Post * Large	-1.26*** (0.356)	-1.30*** (0.359)	1.91*** (0.423)	1.82*** (0.420)	-1.27*** (0.317)	-1.22*** (0.317)
Belgium	-2.27*** (0.217)	2.36*** (0.225)	-7.43*** (0.207)	-4.72*** (0.221)	12.65*** (0.154)	6.21*** (0.140)
Large	-2.66*** (0.702)	-1.20 (0.734)	-3.39*** (0.760)	-2.90*** (0.744)	2.40*** (0.440)	3.11*** (0.466)
Post	7.07*** (0.081)	7.02*** (0.082)	-4.58*** (0.078)	-4.55*** (0.078)	0.80*** (0.056)	0.84*** (0.056)
Dependent variable mean	39.55	39.55	44.11	44.11	11.58	11.58
Sector fixed effects		✓		✓		✓
Baseline controls		✓		✓		✓
Observations	95,318	92,288	95,318	92,288	95,318	92,286

Note: This table presents a replication of the analysis from Tables 1 and 2, using a two-period panel (collapsing the pre- and post-NID periods). The dependent variable in columns 1 and 2 is the equity to total assets ratio, in columns 3 and 4 it is the short-term debt ratio, and in columns 5 and 6 it is the long-term debt ratio. Each specification also contains the same set of baseline control variables as in Tables 1 and 2. Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

A.4 Unbalanced panel

Table A.5: EFFECT OF THE NID ON EQUITY RATIOS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Belgium * Post * Large	5.24*** (0.509)	5.21*** (0.553)	3.27*** (0.241)	2.28*** (0.236)				
Belgium * Post	-1.42*** (0.105)	-0.82*** (0.114)	-0.19** (0.090)	0.12 (0.093)	3.82*** (0.498)	4.38*** (0.541)	3.09*** (0.223)	1.85*** (0.217)
Belgium * Large		-2.14** (0.918)	-0.28 (0.363)	-2.37*** (0.362)				
Post * Large	-2.81*** (0.208)	-2.63*** (0.285)	-2.33*** (0.086)	-1.52*** (0.099)				
Belgium		-0.44*** (0.166)	0.93*** (0.132)	1.86*** (0.148)		-2.58*** (0.903)	0.65* (0.338)	-0.12 (0.346)
Large		-1.23*** (0.388)	-0.99*** (0.127)	0.63*** (0.220)				
Post	7.57*** (0.053)	6.84*** (0.062)	5.95*** (0.040)	-0.70*** (0.053)	4.77*** (0.201)	4.21*** (0.278)	3.62*** (0.076)	-1.57*** (0.103)
Dependent variable mean	38.03	38.03	37.56	37.56	36.42	38.03	37.56	36.69
Small firms	✓	✓	✓	✓				
Group / subsidiary firms			✓	✓			✓	✓
Firm fixed effects	✓				✓			
Sector-year fixed effects				✓				✓
Baseline controls				✓				✓
Observations	1,050,120	1,050,120	2,290,078	1,514,871	45,478	45,478	402,771	271,232

Note: This table presents a replication of the analysis from Table 1, using the larger unbalanced sample. The dependent variable in all columns is the equity to total assets ratio. Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

Table A.6: EFFECT OF THE NID ON DEBT RATIOS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	S/T	S/T	S/T	S/T	S/T	S/T	L/T	L/T
	debt	debt	debt	debt	debt	debt	debt	debt
Belgium * Post * Large	-6.64*** (0.599)	-6.50*** (0.243)	-7.20*** (0.249)				5.40*** (0.181)	
Belgium * Post	1.46*** (0.111)	1.16*** (0.089)	0.97*** (0.092)	-5.18*** (0.589)	-5.34*** (0.226)	-3.91*** (0.230)	-4.02*** (0.074)	0.63*** (0.163)
Belgium * Large	17.87*** (0.960)	14.61*** (0.375)	8.27*** (0.368)				-4.94*** (0.231)	
Post * Large	3.36*** (0.343)	4.73*** (0.107)	5.46*** (0.125)				-3.64*** (0.099)	
Belgium	-6.98*** (0.158)	-6.96*** (0.128)	-6.64*** (0.146)	10.89*** (0.947)	7.64*** (0.353)	1.26*** (0.351)	7.84*** (0.095)	2.78*** (0.213)
Large	-7.82*** (0.433)	-6.93*** (0.147)	-3.46*** (0.222)				2.75*** (0.131)	
Post	-4.72*** (0.061)	-4.37*** (0.042)	-2.99*** (0.075)	-1.36*** (0.338)	0.36*** (0.098)	-2.82*** (0.164)	0.92*** (0.062)	1.68*** (0.134)
Dependent variable mean	45.87	45.46	45.46	45.87	45.46	40.28	11.41	14.10
Small firms	✓	✓	✓				✓	
Group / subsidiary firms		✓	✓		✓	✓	✓	✓
Sector-year fixed effects			✓			✓	✓	✓
Baseline controls			✓			✓	✓	✓
Observations	1,049,433	2,288,652	1,514,519	45,394	402,489	271,161	1,514,233	271,071

Note: This table presents a replication of the analysis from Table 2, using the larger unbalanced sample. The dependent variable in columns 1 to 6 is the short-term debt to total assets ratio and in columns 7 and 8 it is the long-term debt to total assets ratio. Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

A.5 Systematically adding control countries

Table A.7: SYSTEMATICALLY ADDING COUNTRIES

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Belgium * Post * Large	3.24*** (0.265)	3.12*** (0.261)	3.00*** (0.259)	3.00*** (0.259)				
Belgium * Post	0.22** (0.100)	0.15 (0.099)	0.16 (0.099)	0.16 (0.099)	3.46*** (0.245)	3.27*** (0.241)	3.16*** (0.240)	3.16*** (0.240)
Belgium * Large	0.71* (0.425)	0.95** (0.417)	0.84** (0.412)	0.82** (0.411)				
Post * Large	-2.55*** (0.125)	-2.43*** (0.116)	-2.31*** (0.113)	-2.31*** (0.113)				
Belgium	-0.85*** (0.161)	-0.28* (0.160)	-0.26 (0.160)	-0.25 (0.160)	-0.14 (0.393)	0.67* (0.385)	0.58 (0.379)	0.57 (0.379)
Large	-2.02*** (0.221)	-2.27*** (0.204)	-2.15*** (0.193)	-2.13*** (0.193)				
Post	5.87*** (0.053)	5.94*** (0.051)	5.94*** (0.051)	5.93*** (0.051)	3.32*** (0.113)	3.51*** (0.104)	3.62*** (0.100)	3.62*** (0.100)
Sample	Fra	Fra, Ger	Fra, Ger, Neth	All	Fra	Fra, Ger	Fra, Ger, Neth	All
Small firms	✓	✓	✓	✓				
Group / subsidiary firms	✓	✓	✓	✓	✓	✓	✓	✓
Observations	1,071,315	1,127,275	1,152,594	1,154,684	164,776	187,286	204,826	205,966

Note: This table presents results that replicate the analysis from Table 1, systematically adding France, Germany, the Netherlands, and Luxembourg as the control group (compared to Table 1, where all countries are included). The dependent variable in all columns is the equity to total assets ratio for the firm. Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

A.6 Heterogeneous impacts: revenue volatility

In this section, I explore whether the NID had heterogeneous impacts on firms in sectors with the greatest revenue volatility (for whom the performance-contingent nature of equity payments may be particularly beneficial), using the following two specifications:

$$y_{it} = \beta_1 + \beta_2 Post_t + \beta_3 Risky_i + \beta_4 Post_t * Risky_i + \beta_5 \mathbf{X}_{i0} + \varepsilon_{it} \quad (2)$$

$$y_{it} = \beta_1 + \beta_2 Post_t + \beta_3 Large_i + \beta_4 Risky_i + \beta_5 Large_i * Risky_i + \beta_6 Large_i * Post_t + \beta_7 Post_t * Risky_i + \beta_8 Large_i * Post_t * Risky_i + \beta_9 \mathbf{X}_{i0} + \varepsilon_{it} \quad (3)$$

where equation 2 applies to the specification that only includes firms with total assets greater than €4 million, and equation 3 is for the triple-difference specification that includes small firms. *Risky* is a dummy for firms that operate in sectors with above-median volatility of revenues. Specifically, I calculate the pre-NID sector-level coefficient of variation of revenue for all firms, based on their two-digit standard industrial classification (SIC) sector code. Since my dataset contains only two pre-NID years, I construct a sector-level risk measure using Compustat data on US firms using the same two-digit SIC sector code and 13 years of pre-NID data.⁶ \mathbf{X}_{i0} is a matrix of the previously used baseline controls. The main coefficient of interest is β_4 in equation 2 and β_8 in equation 3, which represent the change in leverage in the post-NID period for treatment firms (larger Belgian firms) in sectors with above-median revenue volatility.

Results are presented in Table A.8. Beginning with column 1 of Table A.8, the coefficient on *Post * Large * Risky* of +2.35 indicates that treatment firms in Belgium in sectors with above-median revenue volatility differentially increased their equity ratio post-NID. Column 2 presents results from the same specification as column 1, but replacing the Belgian sample with the French sample. A stark difference is observed; the equivalent French firms (those with assets above €4 million and in sectors with above-median revenue volatility) actually *decreased* their equity ratios in the post-NID period relative to all other firms, with a coefficient on *Post * Large * Risky* of -1.85. Column 3 expands the non-Belgian sample to all surrounding countries, and similarly identifies a negative coefficient on *Post * Large **

⁶ The coefficient of variation is a commonly used measure of risk exposure in finance, insurance and related fields (Brief & Owen, 1969; Fisher, 1959; Hirshleifer, 1988; Kasperski & Holland, 2013; Mahmoudvand & Oliveira, 2018; Osteryoung, Scott, & Roberts, 1977; Rajgopal & Shevlin, 2002; Roberts & Roberts, 1970; Scheel, 1978; Weber, Shafir, & Blais, 2004). The approach of using US data to create a sector-level measure shares similarities with that of Barbiero et al. (2020), who use US sector-level price-to-earnings ratios to proxy for European firms' investment opportunities, and to the approach of Fisman and Love (2007), who argue that the growth rate of US industries is a good proxy for worldwide growth opportunities in countries with developed financial markets. Finally, there are also similarities with Rajan and Zingales (1998), who construct an industry-specific measure of external financial dependence.

Risky of -0.69. At the bottom of the table, *p*-values are reported from a cross-equation test for whether the coefficient on *Post*Large*Risky* differs for Belgian firms compared to the equivalent non-Belgian firms. Unsurprisingly, the null of coefficient equality on *Post*Large*Risky* across the Belgian and non-Belgian specifications is strongly rejected (*p*-values of 0.001 and 0.015 respectively for the test of column 1 compared to columns 2 and 3 respectively). Similar results are observed when using the short-term debt ratio in columns 4 to 6, as well as when restricting the sample by dropping all small firms in columns 7 to 12. The results are also robust to dropping all firms that are part of a larger group / subsidiaries, presented in Table A.9. In Tables A.10 and A.11 I repeat the analysis using the sector-level standard deviation of sales (rather than the coefficient of variation), and results are robust to this alternative specification.

Table A.8: HETEROGENEOUS EFFECTS: RISK EXPOSURE, STAND-ALONE FIRMS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	Equity	Equity	Equity	S/T debt	S/T debt	S/T debt	Equity	Equity	Equity	S/T debt	S/T debt	S/T debt
Post * Large * Risky	2.35* (1.226)	-1.85* (1.022)	-0.69 (0.903)	-2.06* (1.118)	1.02 (1.012)	0.50 (0.905)						
Post * Risky	0.25 (0.258)	0.08 (0.213)	0.10 (0.211)	-1.56*** (0.235)	-0.73*** (0.210)	-0.68*** (0.211)	2.60** (1.182)	-1.76 (1.087)	-0.59 (0.962)	-3.61*** (1.131)	0.29 (1.087)	-0.18 (0.986)
Post * Large	0.90 (0.910)	-0.23 (0.695)	-1.12* (0.619)	-1.87** (0.830)	-0.54 (0.688)	1.66*** (0.621)						
Large * Risky	-6.29*** (1.098)	0.62 (0.914)	0.24 (0.808)	5.64*** (1.001)	1.74* (0.905)	1.55* (0.810)						
Large	2.67*** (0.837)	-1.34** (0.642)	-1.77*** (0.576)	-2.18*** (0.763)	1.21* (0.636)	-0.30 (0.578)						
Risky	-2.15*** (0.267)	-1.53*** (0.205)	-1.67*** (0.203)	2.61*** (0.244)	2.87*** (0.203)	2.42*** (0.204)	-7.94*** (1.253)	1.37 (1.079)	0.22 (0.951)	5.08*** (1.199)	0.18 (1.079)	-0.45 (0.975)
Post	8.32*** (0.246)	10.68*** (0.194)	10.74*** (0.192)	-1.97*** (0.224)	-7.06*** (0.192)	-6.97*** (0.193)	9.98*** (1.240)	10.11*** (1.097)	9.54*** (0.976)	-5.46*** (1.187)	-7.37*** (1.099)	-5.09*** (1.002)
Sample	BEL	FRA	All	BEL	FRA	All	BEL	FRA	All	BEL	FRA	All
Dependent variable mean	38.24	42.07	41.57	40.12	44.18	43.19	38.24	42.07	41.57	40.12	44.18	43.19
Small firms	✓	✓	✓	✓	✓	✓						
Sector fixed effects	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Year fixed effects	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Baseline controls	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Test: Medium*Post*Risky		0.001	0.015		0.017	0.052						
Test: Post*Risky								0.001	0.009		0.002	0.008
Observations	195,420	253,383	263,293	195,419	253,260	263,140	8,440	10,049	13,179	8,440	10,035	13,148

Note: This table presents an analysis of heterogeneous impacts of the NID on firms (both stand-alone and those that part of a larger group / subsidiaries) in sectors with the greatest revenue volatility, measured using a sector-level coefficient of variation of sales (averaged over 13 pre-NID years, based on Compustat data for US firms in the same sector). The variable 'Risky' represents an indicator for a firm being in a sector with an above-median value of the coefficient of variation of sales. The coefficients on Post * Large * Risky (in columns 1 and 4) or Post * Risky (in columns 7 and 10) represent the differential effect of the NID on leverage ratios for the most risk-exposed firms in the treatment group (Belgian large firms). The penultimate two rows of the table display *p*-values for a test whether those coefficients differ for comparably sized firms in control countries (for example, the *p*-values in columns 2 and 3 represent a cross-coefficient test for equality of the coefficient on Post * Large * Risky from those two columns against the coefficient on Post * Large * Risky in row 1). Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** *p*<0.001, ** *p*<0.05, * *p*<0.10.

Table A.9: HETEROGENEOUS EFFECTS: RISK EXPOSURE, ALL FIRMS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	Equity	Equity	Equity	S/T debt	S/T debt	S/T debt	Equity	Equity	Equity	S/T debt	S/T debt	S/T debt
Post * Large * Risky	1.29** (0.558)	0.37 (0.325)	0.36 (0.292)	-0.71 (0.521)	0.18 (0.323)	0.13 (0.299)						
Post * Risky	0.12 (0.216)	-0.07 (0.133)	-0.02 (0.131)	-1.40*** (0.201)	-0.65*** (0.132)	-0.62*** (0.134)	1.41*** (0.524)	0.30 (0.306)	0.34 (0.272)	-2.10*** (0.513)	-0.47 (0.306)	-0.49* (0.283)
Post * Large	-0.01 (0.392)	-2.69*** (0.234)	-2.47*** (0.210)	-1.38*** (0.366)	1.00*** (0.233)	2.43*** (0.215)						
Large * Risky	-2.33*** (0.499)	0.86*** (0.291)	1.19*** (0.261)	2.13*** (0.466)	-1.17*** (0.289)	-0.94*** (0.268)						
Large	0.36 (0.379)	-0.06 (0.226)	-0.33 (0.203)	0.99*** (0.354)	1.36*** (0.224)	0.22 (0.208)						
Risky	-2.93*** (0.217)	-1.52*** (0.129)	-1.76*** (0.126)	4.44*** (0.203)	3.10*** (0.128)	2.28*** (0.129)	-4.84*** (0.521)	-0.44 (0.299)	-0.64** (0.265)	6.60*** (0.510)	1.33*** (0.300)	0.46* (0.275)
Post	8.16*** (0.200)	8.83*** (0.120)	8.94*** (0.117)	-2.43*** (0.187)	-6.76*** (0.119)	-6.45*** (0.120)	9.01*** (0.539)	5.28*** (0.321)	5.76*** (0.285)	-5.37*** (0.528)	-5.68*** (0.321)	-3.63*** (0.296)
Sample	BEL	FRA	All	BEL	FRA	All	BEL	FRA	All	BEL	FRA	All
Dependent variable mean	38.91	41.14	40.39	41.59	47.18	44.88	38.91	41.14	40.39	41.59	47.18	44.88
Small firms	✓	✓	✓	✓	✓	✓						
Sector fixed effects	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Year fixed effects	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Baseline controls	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Test: Medium*Post*Risky		0.087	0.077		0.089	0.113						
Test: Post*Risky								0.025	0.026		0.001	0.001
Observations	337,010	691,898	759,718	337,008	691,711	759,428	48,970	110,367	146,777	48,970	110,332	146,695

Note: This table presents an analysis of heterogeneous impacts of the NID on stand-alone firms in sectors with the greatest revenue volatility, measured using a sector-level coefficient of variation of sales (averaged over 13 pre-NID years, based on Compustat data for US firms in the same sector). The variable 'Risky' represents an indicator for a firm being in a sector with an above-median value of the coefficient of variation of sales. The coefficients on Post * Large * Risky (in columns 1 and 4) or Post * Risky (in columns 7 and 10) represent the differential effect of the NID on leverage ratios for the most risk-exposed firms in the treatment group (Belgian large firms). The penultimate two rows of the table display p-values for a test whether those coefficients differ for comparably sized firms in control countries (for example, the p-values in columns 2 and 3 represent a cross-coefficient test for equality of the coefficient on Post * Large * Risky from those two columns against the coefficient on Post * Large * Risky in row 1). Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

Table A.10: RISK EXPOSURE: STANDARD DEVIATION OF REVENUES, STAND-ALONE FIRMS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	Equity	Equity	Equity	S/T debt	S/T debt	S/T debt	Equity	Equity	Equity	S/T debt	S/T debt	S/T debt
Post * Large * Risky	2.28* (1.229)	-1.82* (1.018)	-0.45 (0.901)	-2.76** (1.122)	1.01 (1.007)	0.05 (0.903)						
Post * Risky	0.32 (0.255)	-0.89*** (0.206)	-0.87*** (0.204)	-1.13*** (0.232)	-0.59*** (0.203)	-0.51** (0.204)	2.60** (1.215)	-2.70** (1.086)	-1.32 (0.962)	-3.89*** (1.161)	0.43 (1.087)	-0.45 (0.986)
Post * Large	0.90 (0.926)	0.01 (0.741)	-1.09* (0.655)	-1.61* (0.845)	-0.63 (0.734)	1.82*** (0.657)						
Large * Risky	-4.75*** (1.101)	0.20 (0.910)	-0.10 (0.806)	5.07*** (1.005)	0.43 (0.901)	1.07 (0.808)						
Large	1.75** (0.851)	-1.22* (0.682)	-1.64*** (0.607)	-1.69** (0.777)	1.79*** (0.675)	-0.15 (0.608)						
Risky	3.25*** (0.283)	0.20 (0.213)	-0.07 (0.210)	-1.69*** (0.258)	3.64*** (0.210)	2.69*** (0.211)	-0.20 (1.332)	2.38** (1.113)	0.21 (0.978)	-0.71 (1.272)	-2.47** (1.114)	-2.24** (1.002)
Post	8.30*** (0.252)	11.07*** (0.198)	11.12*** (0.196)	-2.11*** (0.230)	-7.06*** (0.196)	-6.98*** (0.197)	9.98*** (1.281)	10.77*** (1.132)	9.98*** (1.002)	-5.39*** (1.223)	-7.49*** (1.135)	-4.97*** (1.028)
Sample	BEL	FRA	All	BEL	FRA	All	BEL	FRA	All	BEL	FRA	All
Dependent variable mean	38.24	42.07	41.57	40.12	44.18	43.19	38.24	42.07	41.57	40.12	44.18	43.19
Small firms	✓	✓	✓	✓	✓	✓						
Sector fixed effects	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Year fixed effects	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Baseline controls	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Test: Medium*Post*Risky		0.002	0.030		0.003	0.032						
Test: Post*Risky								0.000	0.001		0.001	0.008
Observations	196,460	255,042	264,962	196,459	254,919	264,809	8,540	10,159	13,289	8,540	10,145	13,258

Note: This table presents an analysis of heterogeneous impacts of the NID on firms (both stand-alone and those that part of a larger group / subsidiaries) in sectors with the greatest revenue volatility, measured using the sector-level standard deviation of sales (averaged over 13 pre-NID years, based on Compustat data for US firms in the same sector). The variable 'Risky' represents an indicator for a firm being in a sector with an above-median value of the standard deviation of sales. The coefficients on Post * Large * Risky (in columns 1 and 4) or Post * Risky (in columns 7 and 10) represent the differential effect of the NID on leverage ratios for the most risk-exposed firms in the treatment group (Belgian large firms). The penultimate two rows of the table display p-values for a test whether those coefficients differ for comparably sized firms in control countries (for example, the p-values in columns 2 and 3 represent a cross-coefficient test for equality of the coefficient on Post * Large * Risky from those two columns against the coefficient on Post * Large * Risky in row 1). Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

Table A.11: RISK EXPOSURE: STANDARD DEVIATION OF REVENUES, ALL FIRMS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	Equity	Equity	Equity	S/T debt	S/T debt	S/T debt	Equity	Equity	Equity	S/T debt	S/T debt	S/T debt
Post * Large * Risky	1.02* (0.558)	0.47 (0.326)	0.55* (0.292)	-0.77 (0.522)	0.42 (0.324)	0.13 (0.299)						
Post * Risky	0.55*** (0.212)	-0.37*** (0.130)	-0.33** (0.129)	-1.01*** (0.199)	-0.64*** (0.130)	-0.59*** (0.132)	1.57*** (0.527)	0.10 (0.308)	0.22 (0.273)	-1.78*** (0.516)	-0.21 (0.308)	-0.46 (0.284)
Post * Large	0.08 (0.407)	-2.72*** (0.241)	-2.55*** (0.216)	-1.40*** (0.380)	0.84*** (0.240)	2.41*** (0.221)						
Large * Risky	-1.40*** (0.500)	-0.14 (0.292)	0.33 (0.262)	1.81*** (0.468)	-1.30*** (0.290)	-0.91*** (0.268)						
Large	0.00 (0.392)	0.42* (0.231)	0.05 (0.208)	1.19*** (0.367)	1.53*** (0.230)	0.27 (0.213)						
Risky	2.81*** (0.228)	-0.21 (0.134)	-0.77*** (0.131)	-1.94*** (0.213)	3.09*** (0.133)	1.65*** (0.134)	0.28 (0.543)	-0.49 (0.314)	-1.09*** (0.275)	1.58*** (0.532)	0.82*** (0.315)	-0.41 (0.286)
Post	7.98*** (0.207)	8.96*** (0.122)	9.07*** (0.119)	-2.53*** (0.194)	-6.73*** (0.121)	-6.43*** (0.122)	8.91*** (0.549)	5.39*** (0.325)	5.83*** (0.288)	-5.48*** (0.538)	-5.82*** (0.326)	-3.64*** (0.300)
Sample	BEL	FRA	All	BEL	FRA	All	BEL	FRA	All	BEL	FRA	All
Dependent variable mean	38.91	41.14	40.39	41.59	47.18	44.88	38.91	41.14	40.39	41.59	47.18	44.88
Small firms	✓	✓	✓	✓	✓	✓						
Sector fixed effects	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Year fixed effects	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Baseline controls	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓	✓
Test: Medium*Post*Risky		0.304	0.369		0.021	0.085						
Test: Post*Risky								0.003	0.005		0.001	0.006
Observations	338,640	694,057	762,137	338,638	693,870	761,847	49,290	110,697	147,337	49,290	110,662	147,255

Note: This table presents an analysis of heterogeneous impacts of the NID on stand-alone firms in sectors with the greatest revenue volatility, measured using the sector-level standard deviation of sales (averaged over 13 pre-NID years, based on Compustat data for US firms in the same sector). The variable 'Risky' represents an indicator for a firm being in a sector with an above-median value of the standard deviation of sales. The coefficients on Post * Large * Risky (in columns 1 and 4) or Post * Risky (in columns 7 and 10) represent the differential effect of the NID on leverage ratios for the most risk-exposed firms in the treatment group (Belgian large firms). The penultimate two rows of the table display p-values for a test whether those coefficients differ for comparably sized firms in control countries (for example, the p-values in columns 2 and 3 represent a cross-coefficient test for equality of the coefficient on Post * Large * Risky from those two columns against the coefficient on Post * Large * Risky in row 1). Standard errors, clustered at the firm level, are reported in parentheses below each coefficient estimate. *** p<0.001, ** p<0.05, * p<0.10.

Finally, Table A.12 provides further details for the 25 sectors that account for 91.5% of firms in the sample, including their value for the coefficient of variation of revenue and the employment growth rate in those sectors between 2000 and 2014 (across the European Union). The data reveal that sectors with above-median revenue volatility are also those that created more jobs (an average increase in employment of 26.8% over the period, compared to 5.0% in the below-median revenue-volatility group). The above-median revenue-volatility group includes sectors in advertising and market research, management consulting and other business services, scientific research and development, as well as information technology and computer programming. In contrast, the below-median revenue-volatility group includes more traditional sectors with slower employment growth (for example, retail trade sectors, printing and publishing, and manufacturing of basic products).

Table A.12: EMPLOYMENT GROWTH RATE BY SECTOR

SECTORS WITH ABOVE-MEDIAN REVENUE VOLATILITY				
SIC CODE	DESCRIPTION	COEFFICIENT OF VARIATION	SAMPLE PROPORTION	EMPLOYMENT GROWTH
50	Wholesale Trade - Durable Goods	5.64	14.2%	12.7%
87	Engineering & Management Services	2.62	10.1%	51.3%
51	Wholesale Trade - Nondurable Goods	2.64	7.2%	12.7%
73	Business Services	7.30	5.6%	37.9%
58	Eating & Drinking Places	2.75	2.7%	39.8%
80	Health Services	3.32	2.3%	28.6%
35	Industrial Machinery & Equipment	4.42	1.7%	1.2%
SIMPLE AVERAGE GROWTH RATE				26.3%
WEIGHTED AVERAGE GROWTH RATE				26.8%

SECTORS WITH BELOW-MEDIAN REVENUE VOLATILITY				
SIC CODE	DESCRIPTION	COEFFICIENT OF VARIATION	SAMPLE PROPORTION	EMPLOYMENT GROWTH
17	Special Trade Contractors	1.95	10.9%	-7.4%
59	Miscellaneous Retail	2.58	8.0%	14.1%
15	General Building Contractors	1.48	4.2%	-7.4%
70	Hotels & Other Lodging Places	2.56	3.3%	39.8%
34	Fabricated Metal Products	2.09	3.1%	-5.7%
20	Food & Kindred Products	2.55	2.3%	6.1%
56	Apparel & Accessory Stores	1.85	2.1%	14.1%
27	Printing & Publishing	1.75	2.0%	-16.4%
75	Auto Repair, Services, & Parking	1.75	1.9%	11.4%
54	Food Stores	1.51	1.8%	14.1%
57	Furniture & Home furnishings Stores	1.85	1.7%	14.1%
24	Lumber & Wood Products	2.40	1.3%	-19.9%
72	Personal Services	1.43	1.2%	36.0%
79	Amusement & Recreation Services	2.00	1.0%	35.9%
52	Building Materials & Gardening Supplies	2.02	0.9%	12.7%
32	Stone, Clay, & Glass Products	1.74	0.8%	-26.2%
55	Automotive Dealers & Service Stations	1.99	0.8%	9.5%
16	Heavy Construction, Except Building	2.05	0.8%	-7.4%
SIMPLE AVERAGE GROWTH RATE				6.5%
WEIGHTED AVERAGE GROWTH RATE				5.0%

Note: The source of employment data is Eurostat (<https://ec.europa.eu/eurostat/cache/metadata/>), which uses Nomenclature of Economic Activities (NACE) codes; these were then mapped onto Standard industrial classification of economic activities (SIC) codes using correspondence tables (<https://ec.europa.eu/eurostat/ramon/relations/>). In cases where a two-digit SIC code mapped onto multiple NACE categories, a further sub-division using three-digit SIC codes was used to map more accurately. For example, SIC code 50 maps onto four NACE categories: (i) advertising and market research; (ii) legal and accounting activities; activities of head offices; management consultancy activities; (iii) architectural and engineering activities; technical testing and analysis, and (iv) scientific research and development. In such cases, a weighted average was used, based on three-digit SIC codes that mapped more directly onto NACE categories. The source for the coefficient of variation of revenues is Compustat, via Wharton Research Data Services (WRDS).

References

- Aghion, P., Besley, T., Brown, J., Caselli, F., Lambert, R., Lomax, R., ... Van Reenen, J. (2013). Investing for prosperity: Skills, infrastructure and innovation. Centre for Economic Performance, London School of Economics and Political Science.
- Barbiero, F., Popov, A., & Wolski, M. (2020). Debt overhang, global growth opportunities, and investment. Journal of Banking & Finance, *120*, 105950.
- Bertrand, M., Duflo, E., & Mullainathan, S. (2004). How much should we trust differences-in-differences estimates? The Quarterly Journal of Economics, *119*(1), 249–275.
- Bilinski, A., & Hatfield, L. A. (2018). Nothing to see here? non-inferiority approaches to parallel trends and other model assumptions. arXiv preprint arXiv:1805.03273.
- Boadway, R., & Bruce, N. (1984). A general proposition on the design of a neutral business tax. Journal of Public Economics, *24*(2), 231–239.
- Boot, A., Carletti, E., Kotz, H., Krahen, J. P., Pelizzon, L., & Subrahmanyam, M. (2020). Coronavirus and financial stability 3.0: Try equity – risk sharing for companies, large and small. VoxEU. doi: <https://voxeu.org/article/try-equity-coronavirus-and-financial-stability>
- Brief, R. P., & Owen, J. (1969). A note on earnings risk and the coefficient of variation. The Journal of Finance, *24*(5), 901–904.
- Célérier, C., Kick, T., & Ongena, S. (2018). Taxing bank leverage: The effects on bank capital structure. Credit Supply and Risk-Taking.
- Cirera, X., Cruz, M., Davies, E., Grover, A., Iacovone, L., Cordova, J. E. L., ... others (2021). Policies to support businesses through the covid-19 shock: A firm level perspective. The World Bank Research Observer, *36*(1), 41–66.
- Devereux, M., & Freeman, H. (1991). A general neutral profits tax. Fiscal Studies, *12*(3), 1–15.
- Economist, T. (2015). A senseless subsidy: Ending the debt addiction. The Economist.
- Fisher, L. (1959). Determinants of risk premiums on corporate bonds. Journal of Political Economy, *67*(3), 217–237.
- Fisman, R., & Love, I. (2007). Financial dependence and growth revisited. Journal of the European Economic Association, *5*(2-3), 470–479.
- Gammie, M. (1991). Equity for companies: A corporation tax for the 1990s. Institute for Fiscal Studies.
- Hall, G. C., Hutchinson, P. J., & Michaelas, N. (2004). Determinants of the capital structures of European SMEs. Journal of Business Finance & Accounting, *31*(5-6), 711–728.

- Hebous, S., & Ruf, M. (2017). Evaluating the effects of ACE systems on multinational debt financing and investment. Journal of Public Economics, *156*, 131–149.
- Hirshleifer, D. (1988). Residual risk, trading costs, and commodity futures risk premia. The Review of Financial Studies, *1*(2), 173–193.
- IMF. (2020). Global financial stability report: Bridge to recovery. International Monetary Fund. doi: "https://www.imf.org/en/Publications/GFSR/Issues/2020/10/13/global-financial-stability-report-october-2020#ExecSum"
- Kahn-Lang, A., & Lang, K. (2020). The promise and pitfalls of differences-in-differences: Reflections on 16 and pregnant and other applications. Journal of Business & Economic Statistics, *38*(3), 613–620.
- Kalemli-Özcan, S., Laeven, L., & Moreno, D. (2018). Debt overhang, rollover risk, and corporate investment: Evidence from the european crisis (Tech. Rep.). National Bureau of Economic Research.
- Kasperski, S., & Holland, D. S. (2013). Income diversification and risk for fishermen. Proceedings of the National Academy of Sciences, *110*(6), 2076–2081.
- Kestens, K., Van Cauwenberge, P., & Christiaens, J. (2012). The effect of the notional interest deduction on the capital structure of Belgian SMEs. Environment and Planning C: Government and Policy, *30*(2), 228–247.
- Klemm, A. (2007). Allowances for corporate equity in practice. CESifo Economic Studies, *53*(2), 229–262.
- Konings, J., Lecocq, C., & Merlevede, B. (2018). Does a tax credit matter for job creation by multinational enterprises? CEPR Discussion Paper No. DP13105.
- Kraus, A., & Litzenberger, R. H. (1973). A state-preference model of optimal financial leverage. The Journal of Finance, *28*(4), 911–922.
- Mahmoudvand, R., & Oliveira, T. A. (2018). On the application of sample coefficient of variation for managing loan portfolio risks. In Recent studies on risk analysis and statistical modeling (pp. 87–97). Springer.
- Mayer, C., Micossi, S., Onado, M., Pagano, M., & Polo, A. (2018). Finance and investment: The European case. Oxford University Press, Oxford.
- McKenzie, D. (2020). Revisiting the difference-in-differences parallel trends assumption: Part I pre-trend testing. World Bank Development Impact. doi: <https://blogs.worldbank.org/impac-tevaluations/revisiting-difference-differences-parallel-trends-assumption-part-ii-what-happen>
- Miller, M. H. (1977). Debt and taxes. The Journal of Finance, *32*(2), 261–275.
- Mirrlees, J., Adam, S., Besley, T., Blundell, R., Bond, S., Chote, R., ... Poterba, J. (2010).

- Reforming the tax system for the 21st century: The Mirrlees Review. Institute for Fiscal Studies.
- Myers, S. C. (1984). The capital structure puzzle. The Journal of Finance, *39*(3), 574–592.
- Myers, S. C., & Majluf, N. S. (1984). Corporate financing and investment decisions when firms have information that investors do not have. Journal of Financial Economics, *13*(2), 187–221.
- OECD. (2020). Insolvency and debt overhang following the COVID-19 outbreak: Assessment of risks and policy responses. OECD Paris. doi: https://www.oecd-ilibrary.org/sites/39a88ab1-en/1/3/2/2/index.html?itemId=/content/publication/39a88ab1-en&_csp_=fd64cf2a9a06f738f45c7aeb5a6f5024&itemIGO=oecd&itemContentType=issue#indicator-d1e6832
- Osteryoung, J. S., Scott, E., & Roberts, G. S. (1977). Selecting capital projects with the coefficient of variation. Financial Management, 65–70.
- Panier, F., Pérez-González, F., & Villanueva, P. (2013). Capital structure and taxes: What happens when you (also) subsidize equity. Working paper.
- Princen, S. (2012). Taxes do affect corporate financing decisions: The case of Belgian ACE. Working paper.
- Quinn, S. (2014). The value of corporate governance for bank finance in an emerging economy: Evidence from a natural experiment. The Journal of Law, Economics, & Organization, *30*(1), 1–38.
- Rajan, R., & Zingales, L. (1995). What do we know about capital structure? some evidence from international data. The Journal of Finance, *50*(5), 1421–1460.
- Rajan, R., & Zingales, L. (1998). Financial development and growth. American Economic Review, *88*(3), 559–586.
- Rajgopal, S., & Shevlin, T. (2002). Empirical evidence on the relation between stock option compensation and risk taking. Journal of Accounting and Economics, *33*(2), 145–171.
- Revoltella, D., Maurin, L., & Pal, R. (2020). Coronavirus and financial stability 3.0: Try equity – risk sharing for companies, large and small. VoxEU. doi: <https://voxeu.org/article/eu-firms-post-covid-19-environment>
- Roberts, C. D., & Roberts, E. N. (1970). Exact determination of earnings risk by the coefficient of variation. The Journal of Finance, *25*(5), 1161–1165.
- Ryan, A. M., Kontopantelis, E., Linden, A., & Burgess Jr, J. F. (2019). Now trending: Coping with non-parallel trends in difference-in-differences analysis. Statistical Methods in Medical Research, *28*(12), 3697–3711.
- Sandbu, M. (2020). The post-pandemic brave new world. International Monetary Fund,

- Finance & Development. doi: <https://www.imf.org/external/pubs/ft/fandd/2020/12/post-pandemic-brave-new-world-sandbu.htm>
- Scheel, W. C. (1978). Comparisons of riskiness as measured by the coefficient of variation. The Journal of Risk and Insurance, *45*(1), 148–152.
- Schepens, G. (2016). Taxes and bank capital structure. Journal of Financial Economics, *120*(3), 585–600.
- Valenduc, C. (2009). Les intérêts notionnels: une réforme fondamentale et controversée. Courrier Hebdomadaire du CRISP(13), 5–52.
- Van Campenhout, G., & Van Caneghem, T. (2013). How did the notional interest deduction affect Belgian SMEs? capital structure? Small Business Economics, *40*(2), 351–373.
- Weber, E. U., Shafir, S., & Blais, A.-R. (2004). Predicting risk sensitivity in humans and lower animals: risk as variance or coefficient of variation. Psychological review, *111*(2), 430.
- Zangari, E. (2014). Addressing the debt bias: a comparison between the Belgian and the Italian ACE systems (Tech. Rep.).